

Research Article

Stolarsky–Puebla index

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Abstract

We introduce a degree-based variable topological index inspired on the Stolarsky mean (known as the generalization of the logarithmic mean). We name this new index as the Stolarsky–Puebla index: $SP_\alpha(G) = \sum_{uv \in E(G)} d_u$, if $d_u = d_v$, and $SP_\alpha(G) = \sum_{uv \in E(G)} [(d_u^\alpha - d_v^\alpha) / (\alpha(d_u - d_v))]^{1/(\alpha-1)}$, otherwise. Here, uv denotes the edge of the network G connecting the vertices u and v , d_u is the degree of the vertex u , and $\alpha \in \mathbb{R} \setminus \{0, 1\}$. We also consider the limiting cases $SP_{\alpha \rightarrow 0}(G)$ and $SP_{\alpha \rightarrow 1}(G)$ that we name as the logarithmic–mean index and the identric–mean index, respectively. Indeed, for given values of α , the Stolarsky–Puebla index reproduces well-known topological indices such as the reciprocal Randić index, the first Zagreb index, and several mean Sombor indices. Moreover, we apply these indices to random networks and demonstrate that $\langle SP_\alpha(G) \rangle$, normalized to the order of the network, scale with the corresponding average degree $\langle d \rangle$. Some mathematical properties of the Stolarsky–Puebla index are also discussed.

Keywords: degree-based topological index; Stolarsky mean; random networks.

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1. Introduction

For two positive real numbers x, y , the Stolarsky mean $S_\alpha(x, y)$ is defined as [22]

$$S_\alpha(x, y) = \lim_{(\xi, \eta) \rightarrow (x, y)} \left(\frac{\xi^\alpha - \eta^\alpha}{\alpha(\xi - \eta)} \right)^{1/(\alpha-1)} = \begin{cases} x & \text{if } x = y, \\ \left(\frac{x^\alpha - y^\alpha}{\alpha(x - y)} \right)^{1/(\alpha-1)} & \text{otherwise,} \end{cases} \quad (1)$$

here, $\alpha \in \mathbb{R} \setminus \{0, 1\}$. In fact, $S_\alpha(x, y)$ is known as the generalization of the logarithmic mean [16]

$$\text{LogMean}(x, y) = \begin{cases} x & \text{if } x = y, \\ \frac{x - y}{\ln x - \ln y} & \text{otherwise.} \end{cases} \quad (2)$$

For given values of α , $S_\alpha(x, y)$ reproduces known means including the logarithmic mean, when $\alpha \rightarrow 0$, and some cases of the power mean [5, 23]

$$PM_\alpha(x, y) = \left(\frac{x^\alpha + y^\alpha}{2} \right)^{1/\alpha}. \quad (3)$$

As examples, in Table 1 we show some expressions for $S_\alpha(x, y)$ for selected values of α with their corresponding names, when available.

Also, there is a well-known inequality relating the Stolarsky mean and the power mean, namely [6, 16, 19]:

$$S_{-1}(x, y) = PM_{\alpha \rightarrow 0}(x, y) \leq S_{\alpha \rightarrow 0}(x, y) \leq PM_{1/3}(x, y) \leq S_2(x, y) = PM_1(x, y) \quad (4)$$

or more explicitly

$$\sqrt{xy} \leq \text{LogMean}(x, y) \leq \left(\frac{x^{1/3} + y^{1/3}}{2} \right)^3 \leq \frac{x + y}{2},$$

where the equality is attained when $x = y$.

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Table 1: Expressions for the Stolarsky mean $S_\alpha(x, y)$ for selected values of α .

| α | $S_\alpha(x, y)$ | name (when available) |
|-----------|---|--|
| $-\infty$ | $S_{\alpha \rightarrow -\infty}(x, y) = \min(x, y)$ | minimum value, $PM_{\alpha \rightarrow -\infty}(x, y)$ |
| -4 | $S_{-4}(x, y) = \left(\frac{x^3 + x^2y + xy^2 + y^3}{4x^4y^4} \right)^{-1/5}$ | |
| -3 | $S_{-3}(x, y) = \left(\frac{x^2 + xy + y^2}{3x^3y^3} \right)^{-1/4}$ | |
| -2 | $S_{-2}(x, y) = \left(\frac{x + y}{2x^2y^2} \right)^{-1/3}$ | |
| -1 | $S_{-1}(x, y) = \sqrt{xy}$ | geometric mean, $PM_{\alpha \rightarrow 0}(x, y)$ |
| 0 | $S_{\alpha \rightarrow 0}(x, y) = \begin{cases} x & \text{if } x = y \\ \frac{x - y}{\ln x - \ln y} & \text{otherwise} \end{cases}$ | LogMean(x, y) |
| $1/2$ | $S_{1/2}(x, y) = \left(\frac{\sqrt{x} + \sqrt{y}}{2} \right)^2$ | $PM_{1/2}(x, y)$ |
| 1 | $S_{\alpha \rightarrow 1}(x, y) = \begin{cases} x & \text{if } x = y \\ \frac{x - y}{x \ln x - y \ln y} & \text{otherwise} \end{cases}$ | identric mean |
| 2 | $S_2(x, y) = \frac{x + y}{2}$ | arithmetic mean, $PM_1(x, y)$ |
| 3 | $S_3(x, y) = \left(\frac{x^2 + xy + y^2}{3} \right)^{1/2}$ | |
| 4 | $S_4(x, y) = \left(\frac{x^3 + x^2y + xy^2 + y^3}{4} \right)^{1/3}$ | |
| ∞ | $S_{\alpha \rightarrow \infty}(x, y) = \max(x, y)$ | maximum value, $PM_{\alpha \rightarrow \infty}(x, y)$ |

2. Stolarsky–Puebla index

A large number of graph invariants of the form

$$TI(G) = \sum_{uv \in E(G)} F(d_u, d_v) \tag{5}$$

are currently being studied in mathematical chemistry; where uv denotes the edge of the graph G connecting the vertices u and v , d_u is the degree of the vertex u , and $F(x, y)$ is an appropriate chosen function, see e.g. [12].

Inspired by the Stolarsky mean and given a simple graph $G = (V(G), E(G))$, here we choose the function $F(x, y)$ in (5) as the Stolarsky mean $S_\alpha(x, y)$ and define the degree–based variable topological index

$$SP_\alpha(G) = \sum_{uv \in E(G)} S_\alpha(d_u, d_v) = \sum_{uv \in E(G)} \begin{cases} d_u & \text{if } d_u = d_v, \\ \left(\frac{d_u^\alpha - d_v^\alpha}{\alpha(d_u - d_v)} \right)^{1/(\alpha-1)} & \text{otherwise,} \end{cases} \tag{6}$$

where uv denotes the edge of the graph G connecting the vertices u and v , d_u is the degree of the vertex u , and $\alpha \in \mathbb{R} \setminus \{0, 1\}$. We name $SP_\alpha(G)$ as the Stolarsky–Puebla index.

Note that for given values of α , $SP_\alpha(G)$ is related to widely studied topological indices: $SP_{-1}(G) = R^{-1}(G)$, where $R^{-1}(G)$ is the reciprocal Randic index [13], $SP_{1/2}(G) = 2^{-2}KA_{1/2,2}^1(G)$, where $KA_{\alpha,\beta}^1(G)$ is the first $(\alpha, \beta) - KA$ index [15], and $SP_2(G) = M_1(G)/2$, where $M_1(G)$ is the first Zagreb index [14]. Also, for selected values of α , $SP_\alpha(G)$ reproduces several mean Sombor indices

$$mSO_\alpha(G) = \sum_{uv \in E(G)} \left(\frac{d_u^\alpha + d_v^\alpha}{2} \right)^{1/\alpha}; \tag{7}$$

recently introduced in [2]. In Table 2, we report some expressions for $SP_\alpha(G)$ for selected values of α that we identify with known topological indices, when applicable.

Table 2: Expressions for the Stolarsky–Puebla index $SP_\alpha(G)$ for selected values of α .

| α | $SP_\alpha(G)$ | index equivalence |
|-----------|--|---|
| $-\infty$ | $SP_{\alpha \rightarrow -\infty}(G) = \sum_{uv \in E(G)} \min(d_u, d_v)$ | $mSO_{\alpha \rightarrow -\infty}(G)$ |
| -4 | $SP_{-4}(G) = \sum_{uv \in E(G)} \left(\frac{d_u^3 + d_u^2 d_v + d_u d_v^2 + d_v^3}{4d_u^4 d_v^4} \right)^{-1/5}$ | |
| -3 | $SP_{-3}(G) = \sum_{uv \in E(G)} \left(\frac{d_u^2 + d_u d_v + d_v^2}{3d_u^3 d_v^3} \right)^{-1/4}$ | |
| -2 | $SP_{-2}(G) = \sum_{uv \in E(G)} \left(\frac{d_u + d_v}{2d_u^2 d_v^2} \right)^{-1/3}$ | |
| -1 | $SP_{-1}(G) = \sum_{uv \in E(G)} \sqrt{d_u d_v}$ | $R^{-1}(G) = mSO_{\alpha \rightarrow 0}(G)$ |
| 0 | $SP_{\alpha \rightarrow 0}(G) = \sum_{uv \in E(G)} \begin{cases} d_u & \text{if } d_u = d_v \\ \frac{d_u - d_v}{\ln d_u - \ln d_v} & \text{otherwise} \end{cases}$ | logarithmic–mean index, see (18) |
| $1/2$ | $SP_{1/2}(G) = \sum_{uv \in E(G)} \left(\frac{\sqrt{d_u} + \sqrt{d_v}}{2} \right)^2$ | $2^{-2} K A_{1/2,2}^1(G) = mSO_{1/2}(G)$ |
| 1 | $SP_{\alpha \rightarrow 1}(G) = \sum_{uv \in E(G)} \begin{cases} d_u & \text{if } d_u = d_v \\ \frac{d_u - d_v}{d_u \ln d_u - d_v \ln d_v} & \text{otherwise} \end{cases}$ | identric–mean index, see (19) |
| 2 | $SP_2(G) = \sum_{uv \in E(G)} \frac{d_u + d_v}{2}$ | $2^{-1} M_1(G) = mSO_1(G)$ |
| 3 | $SP_3(G) = \sum_{uv \in E(G)} \left(\frac{d_u^2 + d_u d_v + d_v^2}{3} \right)^{1/2}$ | |
| 4 | $SP_4(G) = \sum_{uv \in E(G)} \left(\frac{d_u^3 + d_u^2 d_v + d_u d_v^2 + d_v^3}{4} \right)^{1/3}$ | |
| ∞ | $SP_{\alpha \rightarrow \infty}(G) = \sum_{uv \in E(G)} \max(d_u, d_v)$ | $mSO_{\alpha \rightarrow \infty}(G)$ |

3. Computational study of $SP_\alpha(G)$ on random networks

As a first test of the Stolarsky–Puebla index, here we apply it on two models of random networks: Erdős–Rényi (ER) networks and random geometric (RG) graphs. ER networks [8–10, 21] $G_{ER}(n, p)$ are formed by n vertices connected independently with probability $p \in [0, 1]$. While RG graphs [7, 20] $G_{RG}(n, r)$ consist of n vertices uniformly and independently distributed on the unit square, where two vertices are connected by an edge if their Euclidean distance is less or equal than the connection radius $r \in [0, \sqrt{2}]$.

We stress that the computational study of the Stolarsky–Puebla index we perform here is justified by the random nature of the network models we want to explore. Since a given parameter set $[(n, p)$ or $(n, r)]$ represents an infinite-size ensemble of random [ER or RG] networks, the computation of $SP_\alpha(G)$ on a single network is irrelevant. In contrast, the computation of the average value of $SP_\alpha(G)$ on a large ensemble of random networks, all characterized by the same parameter set, may provide useful *average* information about the full ensemble. This *statistical* approach, well known in random matrix theory studies, has been recently applied to random networks by means of topological indices, see e.g. [1, 17, 18]. Moreover, it has been shown that average topological indices may serve as complexity measures equivalent to standard random matrix theory measures [3, 4].

3.1. $SP_\alpha(G)$ on Erdős–Rényi random networks

In what follows, we present the average values of selected Stolarsky–Puebla indices. All averages are computed over ensembles of $10^7/n$ ER networks characterized by the parameter pair (n, p) .

In Figure 1, we present the average Stolarsky–Puebla index $\langle SP_\alpha(G_{ER}) \rangle$ for $\alpha \rightarrow -\infty$, $\alpha \rightarrow 0$, $\alpha \rightarrow 1$, and $\alpha \rightarrow \infty$ as a

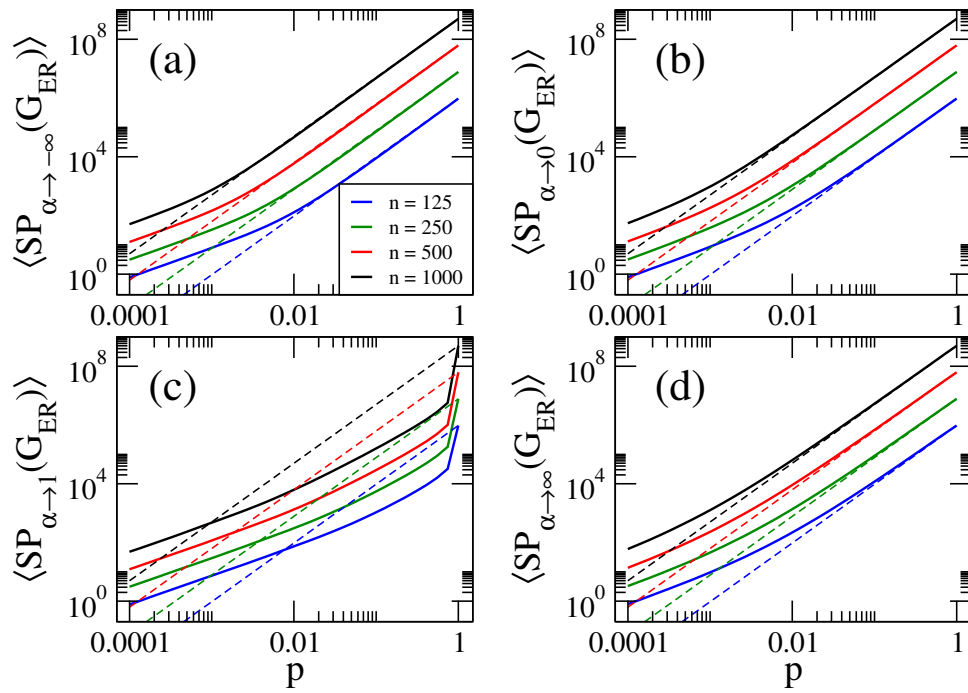


Figure 1: Average value of the Stolarsky–Puebla index $\langle SP_{\alpha}(G_{ER}) \rangle$ as a function of the probability p of Erdős–Rényi networks of size n . Here (a) $\alpha \rightarrow -\infty$, (b) $\alpha \rightarrow 0$, (c) $\alpha \rightarrow 1$, and (d) $\alpha \rightarrow \infty$. Dashed lines correspond to (9).

function of the probability p of ER networks of sizes $n = \{125, 250, 500, 1000\}$. From this figure we observe that the curves of $\langle SP_{\alpha}(G_{ER}) \rangle$ are monotonically increasing functions of p .

We note that in the dense limit, i.e. when $np \gg 1$, we can approximate $d_u \approx d_v \approx \langle d \rangle$ in (6), with

$$\langle d \rangle = (n - 1)p. \tag{8}$$

Thus, when $np \gg 1$, we can approximate $SP_{\alpha}(G_{ER})$ as

$$SP_{\alpha}(G_{ER}) \approx \sum_{uv \in E(G)} d_u \approx \sum_{uv \in E(G)} \langle d \rangle \approx \frac{1}{2}n[(n - 1)p]^2, \tag{9}$$

where we have used $|E(G_{ER})| = n(n - 1)p/2$. In Figure 1, we show that (9) (dashed lines) indeed describes well the data (thick full curves) for large enough p ; except for the case $\langle SP_{\alpha \rightarrow 1}(G_{ER}) \rangle$, see Figure 1(c). We also verified that (9) describes well the data for other values of α , however we did not include them in Figure 1 to avoid figure saturation. We also observed that the smaller the value of α the wider the range of p where the coincidence between (9) and the computational data is observed; compare for example Figs. 1(a) and 1(d), where it is clear that the correspondence of the computational data with (9) is much better in the case of $\alpha \rightarrow -\infty$ than for $\alpha \rightarrow \infty$. In addition, it is relevant to note that (9) does not depend on α .

We also notice that in Figure 1 we present average Stolarsky–Puebla indices as a function of the probability p of ER networks of four different sizes n . It is quite clear from these figures that the curves, characterized by the different network sizes, are very similar but displaced on both axes. This behavior suggests that the average Stolarsky–Puebla indices can be scaled, as will be shown below.

From (9) we observe that $\langle SP_{\alpha}(G_{ER}) \rangle \propto nf[(n - 1)p]$ or

$$\langle SP_{\alpha}(G_{ER}) \rangle \propto nf(\langle d \rangle). \tag{10}$$

Therefore, in Figure 2 we plot again the average Stolarsky–Puebla indices reported in Figure 1, but now normalized to n , as a function of $\langle d \rangle$ showing that all indices are now properly scaled; i.e. the curves painted in different colors for different network sizes fall on top of each other. Moreover, we can rewrite (10) as

$$\frac{\langle SP_{\alpha}(G_{ER}) \rangle}{n} \approx \frac{1}{2} \langle d \rangle^2. \tag{11}$$

In Figure 2, we show that (11) (orange-dashed lines) indeed describe well the computational data (thick full curves) for $\langle d \rangle \geq 10$; except for $\langle SP_{\alpha \rightarrow 1}(G_{ER}) \rangle$, see Figure 2(c).

It is relevant to stress that even when (10) was expected to be valid in the dense limit (i.e. for $\langle d \rangle \gg 1$), it is indeed valid for any $\langle d \rangle$ as clearly seen in Figure 2.

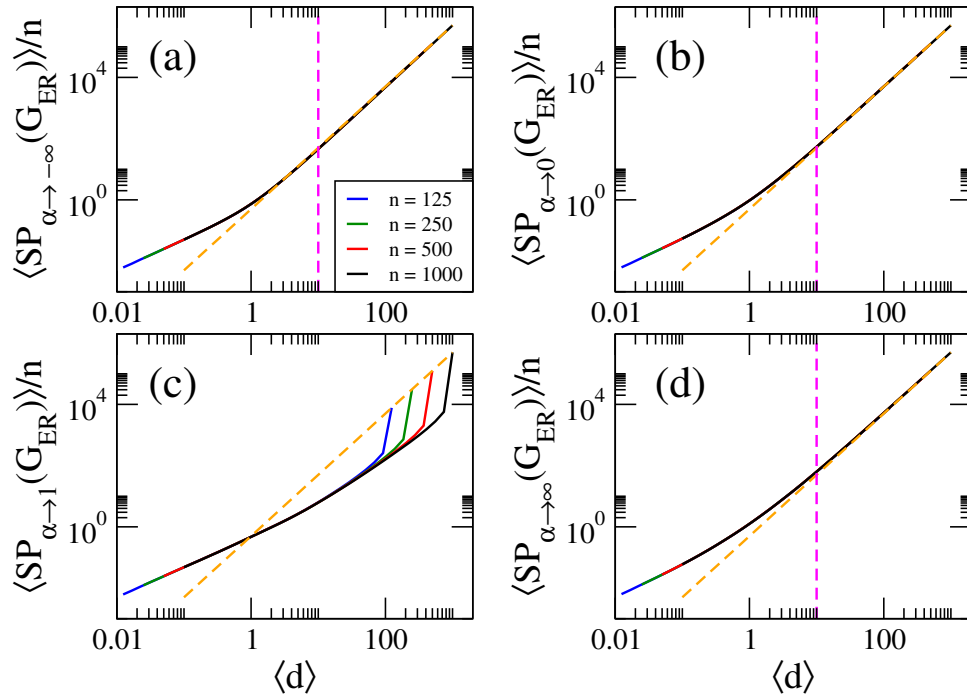


Figure 2: Average value of the Stolarsky–Puebla index $\langle SP_\alpha(G_{ER}) \rangle$, normalized to the network size n , as a function of the average degree $\langle d \rangle$ of Erdős–Rényi networks. Same curves as in Figure 1. Orange dashed lines are (11). The vertical magenta dashed lines indicate $\langle d \rangle = 10$.

3.2. $SP_\alpha(G)$ on random geometric graphs

As in the previous Subsection, here we present the average values of selected Stolarsky–Puebla indices. Again, all averages are computed over ensembles of $10^7/n$ random graphs, each ensemble characterized by a fixed parameter pair (n, r) .

In Figure 3 we present the average Stolarsky–Puebla index $\langle SP_\alpha(G_{RG}) \rangle$ for $\alpha \rightarrow -\infty$, $\alpha \rightarrow 0$, $\alpha \rightarrow 1$, and $\alpha \rightarrow \infty$ as a function of the connection radius r of RG graphs of sizes $n = \{125, 250, 500, 1000\}$. For comparison purposes, Figure 3 is equivalent to Figure 1. In fact, all the observations made in the previous Subsection for ER networks are also valid for RG graphs by just replacing $G_{ER} \rightarrow G_{RG}$ and $p \rightarrow g(r)$, with [11]

$$g(r) = \begin{cases} r^2 \left[\pi - \frac{8}{3}r + \frac{1}{2}r^2 \right] & 0 \leq r \leq 1, \\ \frac{1}{3} - 2r^2 \left[1 - \arcsin(1/r) + \arccos(1/r) \right] + \frac{4}{3}(2r^2 + 1)\sqrt{r^2 - 1} - \frac{1}{2}r^4 & 1 \leq r \leq \sqrt{2}. \end{cases} \quad (12)$$

As well as for ER networks, here, in the dense limit, when $nr \gg 1$, we can approximate $d_u \approx d_v \approx \langle d \rangle$ with

$$\langle d \rangle = (n - 1)g(r). \quad (13)$$

Therefore, in the dense limit, $SP_\alpha(G_{RG})$ is well approximated by:

$$SP_\alpha(G_{RG}) \approx \frac{1}{2}n [(n - 1)g(r)]^2. \quad (14)$$

In Figure 3, we show that (14) (dashed lines) indeed describes well the data (thick full curves) for large enough r ; except for the case $\langle SP_{\alpha \rightarrow 1}(G_{RG}) \rangle$, see Figure 3(c).

It is quite remarkable to note that by substituting the average degree of (13) into (14) we get exactly the same expression of (11):

$$\frac{\langle SP_\alpha(G_{RG}) \rangle}{n} \approx \frac{1}{2} \langle d \rangle^2. \quad (15)$$

So, in Figure 4 we plot again the average Stolarsky–Puebla indices reported in Figure 3 for RG graphs, but now normalized to n , as a function of $\langle d \rangle$ showing that all curves are now properly scaled. Also, in Figure 4, we show that (15) (orange-dashed lines) indeed describes well the computational data (thick full curves) for $\langle d \rangle \geq 10$. We note that as well as for ER networks, here for RG graphs we do not observe the scaling of $\langle SP_{\alpha \rightarrow 1}(G_{RG}) \rangle$.

4. Discussion and conclusions

We have introduced a degree–based variable topological index inspired on the Stolarsky mean, known as the generalization of the logarithmic mean. We named this new index as the Stolarsky–Puebla index $SP_\alpha(G)$, see (6). For given values of α ,

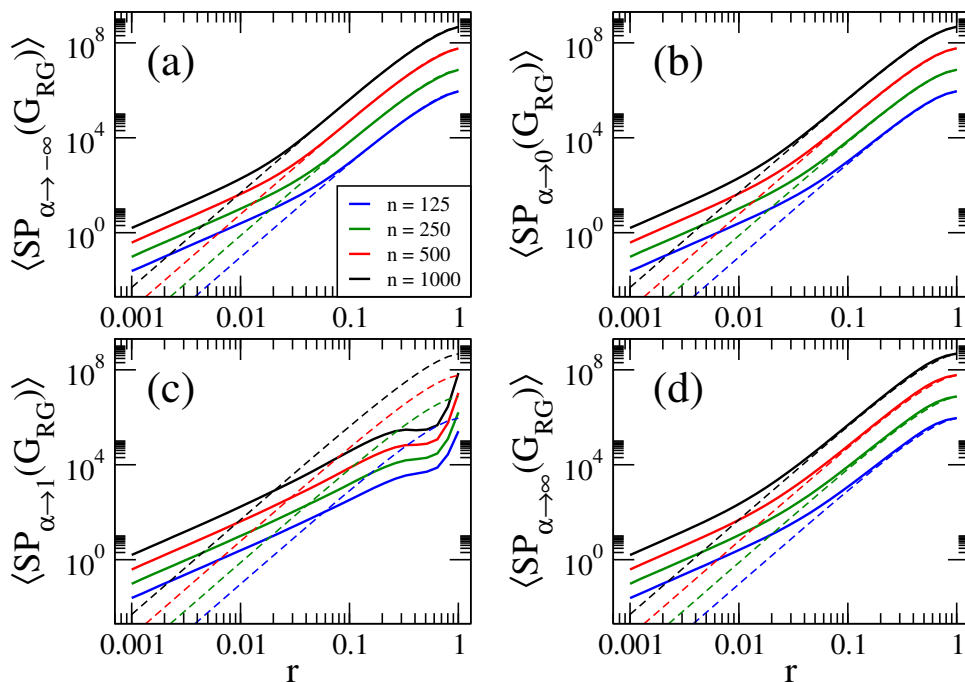


Figure 3: Average value of the Stolarsky–Puebla index $\langle SP_{\alpha}(G_{RG}) \rangle$ as a function of the connection radius r of random geometric graphs of size n . Here (a) $\alpha \rightarrow -\infty$, (b) $\alpha \rightarrow 0$, (c) $\alpha \rightarrow 1$, and (d) $\alpha \rightarrow \infty$. Dashed lines correspond to (14).

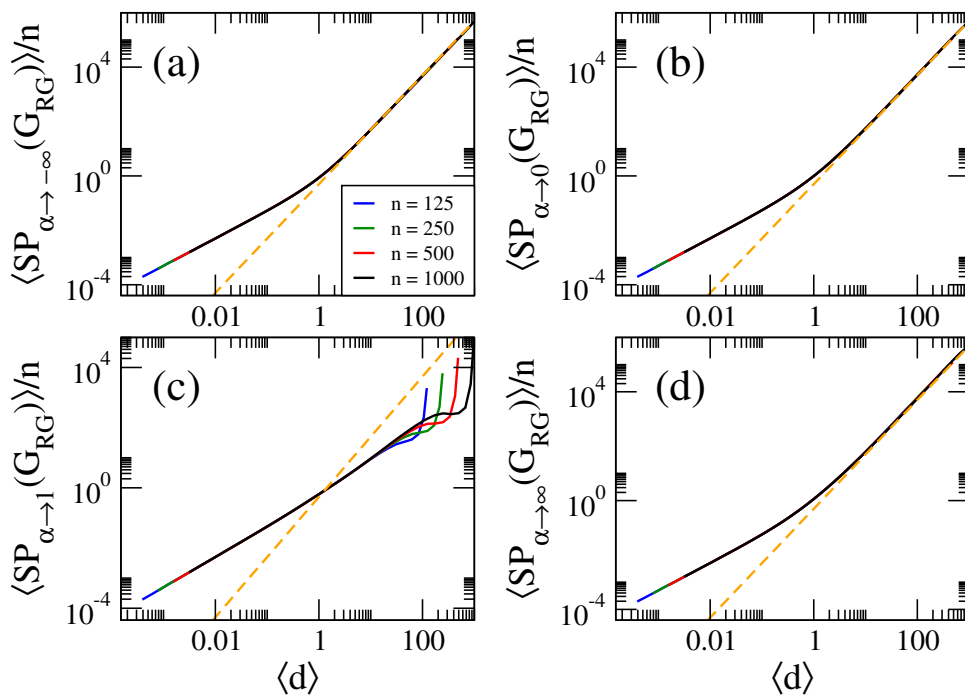


Figure 4: Average value of the Stolarsky–Puebla index $\langle SP_{\alpha}(G_{RG}) \rangle$, normalized to the network size n , as a function of the average degree $\langle d \rangle$ of random geometric graphs. Same curves as in Figure 3. Orange dashed lines are (15). The vertical magenta dashed lines indicate $\langle d \rangle = 10$.

the Stolarsky–Puebla index is related to well-known topological indices, in particular it reproduces several mean Sombor indices $mSO_{\alpha}(G)$, see (7).

We want to add that the inequality of (4) can be straightforwardly used to state inequalities for the indices $SP_{\alpha}(G)$ and $mSO_{\alpha}(G)$, as well as for related indices:

$$SP_{-1}(G) = mSO_{\alpha \rightarrow 0}(G) \leq SP_{\alpha \rightarrow 0}(G) \leq mSO_{1/3}(G) \leq SP_2(G) = mSO_1(G) \tag{16}$$

or

$$R^{-1}(G) \leq \text{LogMean}(G) \leq mSO_{1/3}(G) \leq 2^{-1}M_1(G), \tag{17}$$

which sets bounds for the logarithmic–mean topological index

$$\text{LogMean}(G) = \begin{cases} d_u & \text{if } d_u = d_v, \\ \frac{d_u - d_v}{\ln d_u - \ln d_v} & \text{otherwise,} \end{cases} \tag{18}$$

with respect to the reciprocal Randić index, the mean Sombor index with $\alpha = 1/3$, and the first Zagreb index.

Since there are not many degree–based topological indices including logarithmic functions (as well-known exceptions we can mention the logarithms of the three multiplicative Zagreb indices [12] and the Adriatic indices [24,25]) we want to highlight the release of the logarithmic–mean topological index $\text{LogMean}(G)$ of (18) as well as the identric–mean index

$$\text{idLogMean}(G) = \begin{cases} d_u & \text{if } d_u = d_v, \\ \frac{d_u - d_v}{d_u \ln d_u - d_v \ln d_v} & \text{otherwise,} \end{cases} \tag{19}$$

corresponding to $SP_{\alpha \rightarrow 0}(G)$ and $SP_{\alpha \rightarrow 1}(G)$, respectively.

We have also applied the Stolarsky–Puebla index $SP_{\alpha}(G)$ to Erdős–Rényi (ER) networks and random geometric (RG) graphs and within a statistical random matrix theory approach we demonstrated that $\langle SP_{\alpha}(G) \rangle$, normalized to the order of the network, scales with the corresponding average degree $\langle d \rangle$. However, it is fair to recognize that, for both random network models, $\langle SP_{\alpha \rightarrow 1}(G) \rangle = \langle \text{idLogMean}(G) \rangle$ did not scale; so we believe that the identric–mean index deserves further investigation.

In addition, from (16) we are able to write an equivalent inequality but for the corresponding average values:

$$\langle SP_{-1}(G) \rangle \leq \langle \text{LogMean}(G) \rangle \leq \langle mSO_{1/3}(G) \rangle \leq \langle SP_2(G) \rangle. \tag{20}$$

Indeed, we verified that (20) is satisfied for the both ER random networks and RG graphs (not shown here). Moreover, we computationally found that

$$\langle \text{idLogMean}(G) \rangle \leq \langle SP_{\alpha \neq 1}(G) \rangle, \tag{21}$$

for the two random network models we study here (not explicitly shown here but partially observed in Figs. 1 and 3). The equalities in (20) and (21) are attained when $p = 1$ and $r = \sqrt{2}$, for ER random networks and RG graphs, respectively.

Finally, we want to recall that through a quantitative structure property relationship (QSPR) analysis it was shown [2] that $mSO_{\alpha \rightarrow \pm\infty}(G)$ are good predictors of the standard enthalpy of vaporization, the enthalpy of vaporization, and the heat of vaporization at 25°C of octane isomers. Furthermore, since $SP_{\alpha \rightarrow \pm\infty}(G) = mSO_{\alpha \rightarrow \pm\infty}(G)$, we can conclude that $SP_{\alpha \rightarrow \pm\infty}(G)$ correlate well with the aforementioned physicochemical properties of octane isomers.

In future works we plan to explore mathematical and computational properties of $SP_{\alpha}(G)$, as well as finding optimal bounds and new relationships with known topological indices.

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